Aspirin and fracture risk: a systematic review and exploratory meta-analysis of observational studies

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ABSTRACT

Objectives This review provides insights into the potential for aspirin to preserve bone mineral density (BMD) and reduce fracture risk, building knowledge of the risk-benefit profile of aspirin.

Methods We conducted a systematic review and exploratory meta-analysis of observational studies. Electronic searches of MEDLINE and Embase, and a manual search of bibliographies was undertaken for studies published to 28 March 2018. Studies were included if: participants were men or women aged ≥18 years; the exposure of interest was aspirin; and relative risks, ORs and 95% CIs for the risk of fracture or difference (percentage or absolute) in BMD (measured by dual energy X-ray absorptiometry) between aspirin users and non-users were presented. Risk of bias was assessed using the Joanna Briggs Institute Critical Appraisal Checklists for observational studies. Pooled ORs for any fracture and standardised mean differences (SMDs) for BMD outcomes were calculated using random-effects models.

Results Twelve studies met the inclusion criteria and were included in the meta-analysis. Aspirin use was associated with a 17% lower odds for any fracture (OR 0.83, 95% CI 0.70 to 0.99; I²=71%; six studies; n=511390). Aspirin was associated with a higher total hip BMD for women (SMD 0.03, 95% CI −0.02 to 0.07; I²=0%; three studies; n=9686) and men (SMD 0.06, 95% CI −0.02 to 0.13, I²=0%; two studies; n=4137) although these associations were not significant. Similar results were observed for lumbar spine BMD in women (SMD 0.03, 95% CI −0.03 to 0.09; I²=34%; four studies; n=11330) and men (SMD 0.08; 95% CI −0.01 to 0.18; one study; n=432).

Conclusions While the benefits of reduced fracture risk and higher BMD from aspirin use may be modest for individuals, if confirmed in prospective controlled trials, they may confer a large population benefit given the common use of aspirin in older people.

INTRODUCTION

The anti-inflammatory effects of aspirin via prostaglandin inhibition have recently gained attention. Chronic low-grade inflammation contributes to age-related cardiovascular, neurological, respiratory and musculoskeletal conditions. Low-grade inflammation is associated with increased bone loss and fracture risk. Prostaglandin, an important inflammatory mediator, is likely to have a key role in bone remodelling attributable to inflammation. Prostaglandin E2 stimulates bone resorption and formation and is produced largely from cyclooxygenase-2 induction. Prostaglandins acutely inhibit osteoclast function. However, their chronic effect is to stimulate bone resorption by increasing replication of osteoclast precursors, and differentiation to mature osteoclasts. Aspirin is an inhibitor of prostaglandin production and may influence the cellular basis of bone remodelling responsible for maintaining the material and structural strength of bone. These findings have led to growing interest in the potential for aspirin treatment to provide health benefits broader than cardiovascular, particularly in the prevention of age-related conditions. Osteoporosis—low bone mineral density (BMD) and structural deterioration of bone leading to increased risk of fractures—is the most common bone disease. More than 8.9 million fractures occur annually—equivalent to one fracture
every 3s.7 The combined lifetime risk for hip, forearm and vertebral fractures requiring medical attention is around 40%—equivalent to the risk for cardiovascular disease.8 Despite the availability of several therapies, osteoporosis remains underdiagnosed and undertreated.9 Up to 80% of those with at least one osteoporotic fracture do not receive treatment.9 As the ageing population grows worldwide, including in countries with under-resourced health systems,10–12 so will the demand for therapies that are low-cost, safe and readily available.13 If aspirin use reduces the risk of fracture, the public health impacts may be considerable because even a small benefit in averting or reducing age-related bone fragility may have a large impact at the community level.

There have been no randomised controlled trials (RCTs) reporting on the effects of aspirin treatment and fracture risk in human subjects. While there are increasing numbers of observational studies, findings are contradictory regarding associations between aspirin use and fracture risk.15–26 A review by Chin,27 concluded that aspirin may increase BMD, but its effect on fracture prevention is inconclusive. The limitations of this review were that it included only five observational studies and did not pool quantitative data via meta-analysis or other methods.

Purpose of the current review

Fractures are a growing personal and financial burden to individuals and the healthcare system.7 While there is some indication that aspirin use may be associated with improved BMD and reduced fracture risk, a meta-analysis would yield further insights. Conducting a meta-analysis where data are pooled across several studies can provide an opportunity to identify associations that may have gone undetected in individual studies. The aim of this systematic review was to resolve the controversy in findings from prior studies by conducting an exploratory meta-analysis to examine associations between aspirin use and fracture risk and BMD in adults.28

MATERIALS AND METHODS

Search strategy and selection criteria

We undertook a systematic review and meta-analysis in accordance with the Preferred Reporting Items for Systematic Reviews and Meta-Analyses guidelines. An electronic search for all literature (database inception to 28 March 2018) was performed using Ovid Medline and Embase. The search used MeSH terms (‘aspirin’ OR ‘acetylsalicylic acid’ OR ‘paracetamol’ OR ‘non-steroidal anti-inflammatory drugs’) AND (‘fracture’ OR ‘bone mineral density’ OR ‘osteoporosis’). In addition, a manual search of cited references from retrieved studies, recent systematic reviews, and meta-analyses was conducted (refer online supplementary material I for the list of search terms).

Two authors (JT and Sze-EeSoh) screened titles and abstracts to identify potentially relevant studies, and if a decision could not be made, the full text of the article was examined. Full length articles of selected abstracts were retrieved and assessed independently by two authors (JT and ALB). Studies were considered eligible for the meta-analysis if: (1) participants were men or women aged 18 years or older; (2) the exposure of interest was aspirin; and (3) relative risks, HR or OR and 95% CIs for the risk of fracture or difference (percentage or absolute) in BMD (measured by dual energy X-ray absorptiometry (DXA)) between aspirin users and non-users were presented, or sufficient data were provided to enable these values to be calculated. We included studies that reported aspirin dose and duration either via self-report or identified through prescription histories. Studies were also included if an identifiable and separately analysed subgroup of people using aspirin was presented. Fracture outcomes included the number or rate of any fractures including non-vertebral fractures verified by imaging, while BMD outcomes included measures of bone mass of the whole body as well as at the total hip, proximal femur, femoral neck and spine or lumbar regions. When more than one article presented data from the same study population, the article with more complete reporting of data was selected to avoid double counting participants in the meta-analysis.29

Our aim was to include as many studies as possible in our meta-analysis, considering the limited number of published studies. Any study design was accepted including RCTs, other than single case design studies and systematic reviews. Grey literature such as PhD theses and conference abstracts were excluded to allow a reproducible search strategy. Only papers in English were considered.

Assessment of risk of bias

The risk of bias of included studies was assessed independently by two authors (JT and SES) using the suite of Joanna Briggs Institute (JBI) Critical Appraisal Checklists which have been approved by the JBI Scientific Committee following extensive peer review.30 31 The JBI critical appraisal tools are designed to assess the extent to which a study has addressed the possibility of bias in its design, conduct and analysis.30 These checklists assessed the sample characteristics, ascertainment of exposure, presence of confounding factors, measurement of outcomes, and the use of appropriate statistical techniques. Different criteria were used to appraise methodological quality depending on the study design to minimise bias and improve reliability of findings.32

For cohort studies, the criteria assessed whether groups were similar, if the follow-up time was sufficient to allow outcomes to occur, and loss to follow-up. For case-control studies, the criteria addressed issues related to the definition of cases and controls, selection of control and representativeness of cases, and the comparability of cases and controls. In cross-sectional studies, there was a focus on whether the inclusion criteria and setting of the study were clearly defined. Discrepancies in extracted data and
quality appraisal were discussed and cross-referenced to the original article using a priori decision rules. If consensus could not be gained, agreement was obtained through discussions with a third person (ALB) as per the method described in the Cochrane Handbook for Systematic Reviews of Interventions.33

Data extraction
Data from included studies were independently extracted by two authors (JT and SES) using a customised electronic data extraction form that was piloted with a number of representative studies to ensure the content was appropriate. We extracted information from each study on first author; year and country of study; participant numbers and characteristics; study design; aspirin dose and duration; number of fractures (including site where available); BMD measures for each site (eg, whole body, proximal femur, total hip, femoral neck or lumbar spine); duration of follow-up; and covariates included for adjustment in the analysis.

Statistical analysis
In order to include studies with more than one treatment arm (eg, different aspirin doses) but only one control group, the effect estimate reported for each treatment group was included separately within the meta-analysis. However, the number of participants in the control group were divided equally between the comparisons as per the method described in the Cochrane Handbook for Systematic Review of Interventions.33 This process ensured that control participants were not counted more than once within the meta-analysis.35

Because of the various ways in which the authors reported data on changes in BMD and also wanting to maximise the number of studies and outcomes that could be included in our analysis, we used the standardised mean difference (SMD) approach33 as our summary statistic as opposed to the mean difference (MD) approach. The SMD yields a more conservative estimate than the MD, however, this was considered appropriate given the broad inclusion criteria of this review (which is likely to yield heterogeneity) and exploratory nature of the analysis where we wanted to minimise the likelihood of a type I error (false positive result). This approach allowed a direct comparison of associations across trials that reported either absolute or relative changes in BMD. For each study, the SMD was calculated as follows:

\[
SMD = \frac{\text{mean}_{\text{aspirin}} - \text{mean}_{\text{non-aspirin}}}{\text{SD}_{\text{pooled}}} \left(1 - \frac{3}{4(N_{\text{total}} - 9)}\right) \quad (1)
\]

The MD between aspirin and non-aspirin users was divided by the SD of the difference in BMD for all participants pooled across both treatment groups, and adjusted for small sample bias.33 The SE for the SMD was then calculated using the following formula:

\[
\text{SE}\{\text{SMD}\} = \sqrt{\frac{N_{\text{total}}}{\text{SMD}} + \frac{\text{SMD}^2}{2(N_{\text{total}} - 3.94)}} \quad (2)
\]

If studies lacked SD estimates but provided a SE or CI that related to the MD, we estimated SDs using the following methods33:
1. From SE to SD: the following formula was used:
   \[
   SD = SE \times \sqrt{N}. 
   \]
2. From 95% CI to SE: \( SE = \) \(\frac{(\text{upper limit} - \text{lower limit})}{2 \times 3.92}\)

The OR and 95% CI (if available) was used to describe the association between aspirin use and fracture risk. Effect estimates were manually calculated when needed. Unless otherwise stated, we used the most adjusted risk estimate from each study.

It is known that results may vary systematically with observational study designs.35 As such, for each meta-analysis we performed a subgroup analysis based on study design to explore this. If no significant subgroup differences were identified, study designs were pooled in one-meta-analysis. The meta-analysis included subgroupings for aspirin doses (low dose (≤150mg/day), high dose (>150mg/day) and not reported) and sex in relation to fracture risk and BMD. As sex is an important effect modifier of BMD, results were reported separately for men and women. Sensitivity analyses were also performed by removing each individual study from the fracture and BMD meta-analysis, and removing studies based on study design to assess the robustness of our findings.

Additionally, the effect of aspirin on BMD or fracture risk was examined only when at least five studies reported on the outcome in order to ensure that we had sufficient power to detect a difference between the groups.36 We assessed heterogeneity between studies with the I² statistic as a measure of the proportion of total variation in estimates that is due to heterogeneity. As heterogeneity was suspected in the data due to different study designs, study populations, follow-up duration and aspirin dose and duration, a random effects model was used. If the I² value exceed 75% then a meta-analysis was not undertaken.37 An inverse variance method was also used to weight each estimate. All analyses were conducted with the use of Review Manager, V.5.0 (Revman, The Cochrane Collaboration; Oxford, UK) which allowed estimates and their SE to be entered directly using the ‘Generic Inverse Variance’ outcome to facilitate the analysis of continuous and dichotomous outcome data from non-randomised studies.33 38

Patient and public involvement
There was no direct patient or public involvement in this review.

RESULTS
Description of studies
There were 655 articles identified, 12 were duplicates and 601 were excluded after reviewing the titles and abstracts, leaving 46 articles. An additional four articles were found from reviewing references from articles obtained in full text. After full text review, a total of 14 articles were assessed as eligible for inclusion. However,
only 12 studies were used as three reported different outcomes on the same sample (figure 1). To avoid double counting, only data from Vestergaard et al was used for the meta-analysis as it reported the most complete dataset. Of the 12 unique studies, three were prospective cohort studies, three were case-control studies, and six were cross-sectional studies (table 1). Four studies reported fracture outcomes only, six studies reported BMD outcomes only, and the remaining four studies reported both fracture and BMD outcomes (online supplementary materials 1 and 2). We were unable to identify any RCTs of aspirin that included data on fracture or BMD outcomes.

**Risk of bias in included studies**

The risk of bias assessment is reported in table 2. The risk of bias was considered low for the three cohort studies with respect to the exposure, identifying and dealing with confounders, measuring outcomes, having sufficient follow-up time and using appropriate statistical analysis. Measurement of fracture and BMD outcomes were considered reliable and valid. Fractures were verified using radiological reports or obtained using the International Classification of Diseases 10 codes (online supplementary material 2), while all BMD outcomes were measured using DXA scans (online supplementary material 3). In contrast, reasons for loss of follow-up were not well documented, with only one describing strategies to address incomplete follow-up. For the case-control studies, the risk of bias was low for most studies except for one which did not provide sufficient information regarding the cases and controls, as well as how exposure to aspirin was measured. The risk of bias for identifying confounders, measuring outcomes and using appropriate statistical analysis was low for all cross-sectional studies. However, inclusion criteria, study subjects, and measurement of exposure were poorly defined in a number of studies.

**Meta-analysis**

Effect estimates (ie, ORs) for the association between aspirin use and fracture risk were manually calculated for two studies. The MD in BMD outcomes were reported by three studies, with the remaining studies reporting percentage difference in BMD. Using equations (1) and (2), the SMD and SE for the SMD was calculated for each study and used in the meta-analysis.

No significant differences were identified when we performed a subgroup analysis based on study design (figure 2). Thus, data from included studies were pooled into one meta-analysis for fracture and BMD outcomes. Data pooled from six studies that included 511390 people indicated aspirin was associated with a 17% lower odds of any fracture (OR 0.83, 95% CI 0.70 to 0.99; I²=70%; figure 3). There was no significant difference in

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**Figure 1** Preferred Reporting Items for Systematic Reviews and Meta-Analyses flowchart of study selection process.
Table 1 Characteristics of included studies

<table>
<thead>
<tr>
<th>Study ID</th>
<th>Study design</th>
<th>Population</th>
<th>Sample size</th>
<th>Source of participants</th>
<th>Age, mean</th>
<th>Female, %</th>
<th>Follow-up (years)</th>
<th>Exposure to aspirin</th>
<th>Outcomes</th>
</tr>
</thead>
<tbody>
<tr>
<td>Bauer et al (1996)</td>
<td>PC</td>
<td>USA</td>
<td>7786</td>
<td>Community*</td>
<td>73.1</td>
<td>100</td>
<td>1.6</td>
<td>Self-report</td>
<td>1–4 times/week</td>
</tr>
<tr>
<td>Bleicher et al (2011)</td>
<td>CS</td>
<td>Australia</td>
<td>1705</td>
<td>Community</td>
<td>77.0</td>
<td>0</td>
<td>–</td>
<td>Medication verified in clinic†</td>
<td>NR</td>
</tr>
<tr>
<td>Bonten et al (2017)</td>
<td>CS</td>
<td>Netherlands</td>
<td>854</td>
<td>Community</td>
<td>59.0‡</td>
<td>34‡</td>
<td>–</td>
<td>Medication verified in clinic†</td>
<td>30–125 mg/day</td>
</tr>
<tr>
<td>Carbone et al (2003)</td>
<td>CS</td>
<td>USA</td>
<td>2853</td>
<td>Community</td>
<td>73.6</td>
<td>50</td>
<td>–</td>
<td>Medication verified in clinic†</td>
<td>328 mg/day</td>
</tr>
<tr>
<td>Chuang et al (2016)</td>
<td>CC</td>
<td>Taiwan</td>
<td>555</td>
<td>Community</td>
<td>74.0§</td>
<td>61§</td>
<td>5</td>
<td>Prescription history</td>
<td>106 mg§</td>
</tr>
<tr>
<td>Dobnig et al (2007)</td>
<td>PC</td>
<td>Austria</td>
<td>1664</td>
<td>Nursing homes</td>
<td>–</td>
<td>100</td>
<td>2</td>
<td>NR</td>
<td>NR</td>
</tr>
<tr>
<td>Hill et al (2008)</td>
<td>CS</td>
<td>Trinidad and Tobago</td>
<td>340</td>
<td>Community¶</td>
<td>63.9</td>
<td>100¶</td>
<td>–</td>
<td>Medication verified in clinic†</td>
<td>≥3 times/week</td>
</tr>
<tr>
<td>Hill et al (2008)</td>
<td>CS</td>
<td>Trinidad and Tobago</td>
<td>2501</td>
<td>Community</td>
<td>56.3</td>
<td>0</td>
<td>–</td>
<td>Self-report</td>
<td>NR</td>
</tr>
<tr>
<td>Lane et al (1997)</td>
<td>CS</td>
<td>USA</td>
<td>499</td>
<td>Community*</td>
<td>73.6</td>
<td>100</td>
<td>–</td>
<td>Self-report</td>
<td>5–7 days/week</td>
</tr>
<tr>
<td>Vestergaard et al (2006, 2012)</td>
<td>CC</td>
<td>Denmark</td>
<td>498617</td>
<td>Community</td>
<td>43.4</td>
<td>52</td>
<td>1</td>
<td>Prescription history</td>
<td>≤150 mg/day</td>
</tr>
<tr>
<td>Vestergaard et al (2012)</td>
<td>PC</td>
<td>Denmark</td>
<td>2016</td>
<td>Community*¶</td>
<td>50.8‡</td>
<td>100</td>
<td>10</td>
<td>Self-report</td>
<td>325 mg/day</td>
</tr>
<tr>
<td>Williams et al (2011)</td>
<td>Nested CC</td>
<td>Australia</td>
<td>1344</td>
<td>Community</td>
<td>–</td>
<td>100</td>
<td>2</td>
<td>Medication verified in clinic†</td>
<td>NR</td>
</tr>
</tbody>
</table>

*Causasian women only.
†Participants were asked to bring all prescription and non-prescription medication for verification.
‡Age and percentage of females for aspirin uses only.
§Age, percentage of females and aspirin dosage for cases only.
¶Postmenopausal women only.
BMD, bone mineral density; CC, case-control; CS, cross-sectional; PC, prospective cohort.
Table 2  Risk of bias assessment using the Joanna Briggs Institute critical appraisal Checklists<sup>30</sup>

<table>
<thead>
<tr>
<th>Sample characteristics</th>
<th>Exposure</th>
<th>Confounding</th>
<th>Outcomes and analysis</th>
</tr>
</thead>
<tbody>
<tr>
<td></td>
<td></td>
<td></td>
<td>Groups free of outcome of exposure</td>
</tr>
<tr>
<td>Similar groups</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Comparable groups</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Matched cases and controls</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Same criteria for identification of controls</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Inclusion criteria defined</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Participants and setting described</td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Measurement of exposure similar for cases and controls</td>
<td>Yes</td>
<td>Yes</td>
<td>–</td>
</tr>
<tr>
<td>Valid and reliable measure of exposure</td>
<td>Yes</td>
<td>Yes</td>
<td>–</td>
</tr>
<tr>
<td>Objective, valid criteria to measure condition</td>
<td>Yes</td>
<td>Yes</td>
<td>–</td>
</tr>
</tbody>
</table>

**Cohort studies**

- **Bauer et al (1996)<sup>15</sup>**
  - No – – – – – Yes Yes – Yes Yes UC Yes No – Yes
- **Dobney et al (2007)<sup>16</sup>**
  - Yes – – – – – Yes No – Yes Yes Yes Yes Yes No Yes – Yes
- **Vestergaard et al (2012)<sup>22</sup>**
  - Yes – – – – – Yes Yes – Yes Yes Yes Yes Yes UC No – Yes

**Case-control studies**

- **Chuang et al (2016)<sup>25</sup>**
  - – Yes Yes Yes Yes Yes Yes – Yes Yes No* Yes No* No* Yes
- **Vestergaard et al (2006, 2012)<sup>19 21 26</sup>**
  - Yes Yes Yes Yes Yes – Yes Yes Yes Yes Yes Yes UC No – Yes
- **Williams et al (2011)<sup>20</sup>**
  - – Yes No UC – – Yes Yes – UC UC – Yes – – – Yes No*

**Cross-sectional studies**

- **Bleicher et al (2011)<sup>14</sup>**
  - – – – – No Yes – – Yes No Yes UC Yes – – – Yes
- **Bonten et al (2017)<sup>14</sup>**
  - – – – – Yes Yes – – Yes Yes Yes Yes Yes – Yes – – – Yes
- **Carbone et al (2003)<sup>14</sup>**
  - – – – – Yes No – – Yes Yes Yes Yes – Yes – – – Yes
- **Hill et al (2008)<sup>17</sup>**
  - – – – – Yes Yes – – Yes Yes Yes Yes – Yes – – – Yes
- **Hill et al (2008)<sup>20</sup>**
  - – – – – Yes Yes – No No Yes Yes Yes Yes – Yes – – – Yes
- **Lane et al (1997)<sup>23</sup>**
  - – – – – No No – Yes Yes Yes Yes UC Yes – Yes – – – Yes

*The statistical analysis undertaken for the primary analysis was adjusted; however, for the purposes of this review, effect estimates for association between aspirin use and fracture risk were manually calculated and no adjustment for covariates was made.

LTFU, Lost to follow-up; No, criteria not met; UC, unclear; Yes, criteria met.
Figure 2 Meta-analysis of the association by study design of the association between aspirin use and (A) risk of any fracture (adjusted OR); (B) hip bone mineral density; and (C) lumbar spine bone mineral density.
associations across dose subgroups (figure 3). A subgroup analysis on studies that included only women was not undertaken as considerable heterogeneity was observed with an I² value of 78%. There were no studies that examined risk of fractures only in men.

Aspirin was associated with a higher total hip BMD for women (SMD 0.03, 95% CI −0.02 to 0.07; I²=0%; three studies; n=9686; figure 4A) and men (SMD 0.06, 95% CI −0.02 to 0.13, I²=0%; two studies; n=4137; figure 4A) although these associations were not significant. There were also no significant differences in associations across dose (figure 4B). Similar results were observed for lumbar spine BMD in women (SMD 0.03, 95% CI −0.03 to 0.09; I²=34%; four studies; n=11390; figure 4C) and men (SMD 0.08, 95% CI −0.01 to 0.18; one study; n=432; figure 4C).

Sensitivity analysis was performed whereby each study was removed from the fracture and BMD meta-analysis. The direction of association remained when individual studies were removed but significance often changed. For example, when the Williams et al study was removed from the analysis, aspirin was associated with a 7% lower odds of any fracture (OR 0.93, 95% CI 0.91 to 0.95). When we removed the Bauer et al study from the lumbar spine BMD meta-analysis, aspirin use was associated with higher spine BMD in women (SMD 0.04, 95% CI 0.10 to 0.17) and men (SMD 0.08, 95% CI −0.01 to 0.18). Similar results were obtained when we removed individual studies based on type of design.

No meta-analyses were undertaken for femoral neck or total body BMD as there were too few studies. Only four studies reported on femoral neck BMD16 17 22 23 and two reported on total body BMD with conflicting results.14 22 Two studies reported no association between aspirin and femoral neck BMD,22 23 while two reported a higher femoral neck BMD in aspirin users compared with non-users.16 17 Likewise, only one of the studies that reported on total body BMD reported a higher total BMD in aspirin users compared with non-users.14

**DISCUSSION**

**Main findings**

Aspirin is one of the most commonly used medications worldwide and has the potential to benefit and/or prevent many chronic diseases. This review extends previous reviews of aspirin and bone27 by including additional studies16–18 20 23–25 40 and an exploratory meta-analysis. While firm conclusions about causation cannot be inferred, and nor was this the intent of our exploratory review, our findings suggest that aspirin use was associated with a lower risk of any fracture and potentially higher BMD at the spine and hip. These findings are consistent with preclinical and animal studies.27 While the reduction in fracture risk was small, given the widespread use of aspirin and the high prevalence of undetected low bone density, even small reductions in fracture risk may yield considerable benefits at a population level. Even though we do not expect that aspirin would be considered as a primary fracture prevention therapy, this study builds on the knowledge of the potential broader benefits of aspirin therapy for those taking it for cardiovascular indications.

The consistent findings of reduced risk of fracture across studies included in this review is encouraging. It is important to keep in mind that studies were quite diverse in design, populations included, data collection methods and follow-up periods and we did observe high heterogeneity especially for fracture risk. While we need to interpret this finding with some caution, there appeared to be a consistent indication that aspirin use is associated with positive bone outcomes. While our findings suggest that dose and

<table>
<thead>
<tr>
<th>Study or Subgroup</th>
<th>Aspirin users</th>
<th>Non-aspirin users</th>
<th>Odds Ratio IV, Random, 95% CI</th>
<th>Odds Ratio IV, Random, 95% CI</th>
</tr>
</thead>
<tbody>
<tr>
<td>1.1.1 Low dose aspirin (≤51mg/day)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Chueung (2016) (25)</td>
<td>-0.046, 0.38</td>
<td>55</td>
<td>500</td>
<td>4.7%</td>
</tr>
<tr>
<td>Vestergaard (2012) (≤51mg/day) (26)</td>
<td>-0.073, 0.01</td>
<td>37879</td>
<td>460738</td>
<td>29.3%</td>
</tr>
<tr>
<td>Subtotal (95%, CI)</td>
<td>37934</td>
<td>461238</td>
<td>33.9%</td>
<td></td>
</tr>
<tr>
<td>Heterogeneity: Tau²=0.00, Chi²=0.01, df=1 (P=0.94), I²=0%</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
<tr>
<td>Test for overall effect: Z=7.30 (P = 0.00001)</td>
<td></td>
<td></td>
<td></td>
<td></td>
</tr>
</tbody>
</table>

| 1.1.2 High dose aspirin (>150mg/day) | | | | |
| Vestergaard (2012) 325±26mg/day (22) | -0.062, 0.18 | 268 | 1748 | 13.4% | 0.94 [0.66, 1.34] |
| Subtotal (95%, CI) | 268 | 1748 | 13.4% |
| Heterogeneity: Not applicable |
| Test for overall effect: Z=0.34 (P = 0.73) |

| 1.1.3 Dose not reported (mg/day) | | | | |
| Bauer (1990) 5-7 times/week (15) | 0.1 | 1413 | 5781 | 21.5% | 1.00 [0.82, 1.22] |
| Dabng (2007) (18) | -0.329, 0.16 | 472 | 1152 | 15.1% | 0.72 [0.53, 0.96] |
| Williams (2011) (20) | -0.631, 0.15 | 241 | 1103 | 16.1% | 0.53 [0.40, 0.71] |
| Subtotal (95%, CI) | 2126 | 8076 | 52.7% |
| Heterogeneity: Tau²=0.10, Chi²= 12.81, df=2 (P=0.002),I²=84% |
| Test for overall effect: Z=1.56 (P = 0.12) |

| Total (95%, CI) | 40328 | 471062 | 100% | 0.83 [0.70, 0.99] |
| Heterogeneity: Tau²=0.03, Chi²= 18.87, df=5 (P=0.005),I²=70% |
| Test for overall effect: Z=22.05 (P = 0.04) |
| Test for subgroup differences: Chi²= 1.41, df=2 (P=0.49),I²=0% |

**Figure 3** Meta-analysis of the association between aspirin use and risk of any fracture (adjusted OR) by aspirin dosage.
Figure 4  Meta-analysis of the association between aspirin use and bone mineral density in the (A) hip by sex; (B) hip by aspirin dosage; (C) lumbar spine by sex; and (D) lumbar spine by aspirin dosage.
sex subgroups do not influence bone outcomes, these subgroup analyses were likely underpowered. Only eight of the 14 publications reported aspirin dosage in terms of frequency and dose. There were also no studies that reported on aspirin use and fracture outcomes in men only. In contrast, six of the studies were conducted in women only. Future studies should aim to collect and report findings for different aspirin doses and sex to enable dose-and-or sex related effects (if any) to be ascertained.

When low-dose aspirin use is recommended for prevention of cardiovascular disease by a healthcare professional, nine of 10 adults will follow the recommendation. This suggests the drug is broadly accepted by patients as a treatment regime. The findings of this review build on the existing evidence that supports the role of aspirin in the prevention of conditions including cancer and depression. As the population ages, increasing numbers of older people will be taking aspirin. Knowledge of the relative benefits and risks beyond cardiovascular in older individuals is critical to inform treatment decisions. If future RCTs provide evidence of positive effects on bone and fracture prevention, it will be reassuring for those taking aspirin for cardiovascular indications to know that the drug has additional positive health impacts beyond those for which its primary therapeutic effect was intended.

Limitations
Some caution is needed in interpreting the findings. The search strategy included full-text reports published in English and we may have missed studies in other languages. In addition, meta-analyses of observational studies are subject to the same biases and confounding in the original studies such as confounding by indication and unmeasured factors, self-report of aspirin use, and survival bias. We minimised these limitations by: (1) developing explicit and broad criteria that focused on extracting the best available evidence; (2) utilising a broad search strategy to ensure that all relevant literature was included; and (3) exploring sources of heterogeneity by conducting subgroup analyses, none of which revealed any significant differences in findings. Importantly, several studies reported aspirin users and non-users differed at baseline in risk factors for fracture. These included: age; sex; race; weight; oestrogen therapy; prescription of medications known to increase falls risk; higher levels of comorbidity; number of prior fractures; and existing diagnosis of osteoporosis, osteoarthritis or rheumatoid arthritis. While analyses often adjusted for these factors, the risk that groups were systematically different in factors neither measured nor adjusted for, remains high. For example, no study reported adjustment for physical activity. Aspirin use could also be a marker for an unmeasured confounder. This confounding by indication, and resulting imbalance in the underlying risk profile between aspirin users and non-users, may have biased results. Further, data on confounding variables were often self-reported. Most studies were subject to recall bias particularly with respect to use of aspirin. All studies relied on participant self-reporting of current and prior aspirin use over long periods of time (range, 1–10 years), which is known to over-estimate actual drug use, with two studies being completed retrospectively. This is likely to increase the potential effect of aspirin on fracture risk and BMD. Finally, none of the studies reported fracture or BMD results separately for first or subsequent fractures. The cause of fractures (eg, trauma) were also not reported by included studies. As such, we are unable to determine whether the potential protective association of aspirin is for primary or secondary prevention. This would be valuable to explore in future research.

Implications
The WHO has identified fracture prevention as a public health priority. By the year 2040, the number of people at high risk of sustaining a fracture will double to 319 million. The evidence presented in this review provides a window into the potential for aspirin to be a novel therapy for fracture prevention. While causation cannot be inferred from our review findings, we have provided a basis to inform future research including RCTs that aim to examine the effect of aspirin on fracture risk. Given that aspirin use is not without risk any potential skeletal benefits must be considered in the broader context of therapeutic risk. Studies in older people report long-term use of aspirin is associated with increased risk of major bleeding including haemorrhagic stroke. Currently there is an absence of RCT data in those aged 70 years and older who have an increased susceptibility to adverse drug effects due to impaired hepatic and renal function, polypharmacy and comorbidity. Thus, there is a need for future research, particularly RCTs in older people to focus on assessing data on multiple end-points (such as cardiovascular, bleeding, fractures and mortality) in the one population to provide accurate estimates of the relative benefits and risks of using aspirin.

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